

Recognition of Short-Time Specific Random Elements in Random Sequences

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Received: June 2010; accepted: January 2011

Abstract. In this paper we consider random sequences in the background of which specific short-time random elements can emerge. The theory and constructive methods for recognition of short-time specific random elements that may emerge in the background of random sequences are expounded. The results of experimental investigations are presented. The prospects for a wider application of the results obtained are discussed as well.

Keywords: recognition, random sequences, short-time random events, stochastic dynamic systems.

1. Introduction

The operation of dynamic systems is accompanied by random phenomena of a varying nature that reflect their functioning peculiarities. By analyzing the signals that describe the properties of these stochastic phenomena, one can recognize the states of dynamic systems (Atanasov *et al.*, 2010; Pupeikis, 2010). We can find out whether the operation mode is normal or of pre-emergency state, dangerous and trustful. We can also measure the time instants when the properties of dynamic systems are changing (Willsky, 1976; Shaban, 1980; Kassam, 1980; Kligienė *et al.*, 1983; Basseville *et al.*, 1985, 1993; Telksnys, 1986, 1987). Here we face another set of problems, when short-time specific random events emerge at random time moments, while dynamic system is operating. Then, in the background of a random signal that characterizes normal operation of the dynamic system, we have to recognize the short-time random signals appearing at random time moments. Problems of this nature arise, for instance, when solving the cardiovascular system's functional condition evaluation problems (Klersy *et al.*, 2009; Goss *et al.*, 2009). With an intention to solve the problems of this kind, we further present the recognition theory and a constructive method of short-time specific random signals that appear at random time moments in the background of a random signal.

2. Statement of the Problem

Let us consider a random sequence

$$Y(i) = X(i) + S(i), \quad i = 1, \dots \quad (1)$$

The component $X(i)$, $i = 1, \dots$ in it is a random sequence described by the Gauss law with the mean $M = EX(i)$, $i = 1, \dots$, standard deviation $D = [E(X(i) - M)^2]^{1/2}$, $i = 1, \dots$ and covariance function $K(j) = E\{[X(i) - M][X(i + j) - M]\}$, $i = 1, \dots, j = 1, \dots, J$. $M, D, K(j)$, $j = 1, \dots, J$ – are unknown characteristics.

The second component is represented by an expression

$$S(i) = \begin{cases} C(i), & i = i1, i2, \dots, iL, \\ 0, & i \neq i1, i2, \dots, iL, \end{cases} \quad (2)$$

where $C(i)$, $i = i1, i2, \dots, iL$ short time specific random elements are elements of a single sequence of random amplitude that emerge of random time moments.

We observe the sample

$$y(i) = x(i) + s(i), \quad i = 1, \dots, N \quad (3)$$

of a random sequence $Y(i) = X(i) + S(i)$, $i = 1, \dots$

In (3) $x(i)$, $i = 1, \dots, N$ is a sample of the random sequence $X(i)$, $i = 1, \dots$ and

$$s(i) = \begin{cases} c(i), & i = i1, i2, \dots, iL, \\ 0, & i \neq i1, i2, \dots, iL, \end{cases}$$

are a sample of the random sequence $S(i)$.

We need to determine the argument values $i = i1, i2, \dots, iL$ of appearance of the short-time random specific elements $y(i) = x(i) + s(i)$, $s(i) \neq 0$, $i = i1, i2, \dots, iL$.

3. Solving of the Problem

First, we analyze situation when

$$Y(i) = X(i) + S(i), \quad i = 1, \dots, \\ S(i) = \begin{cases} C(i), & i = i1, \\ 0, & i \neq i1. \end{cases} \quad (4)$$

Define a sequence of random variables

$$U(i) = [X(i) - M]/D, \quad i = 1, 2, \dots, \quad (5)$$

M and D are unknown.

Instead of M and D we use their estimates m and d :

$$m = \frac{1}{N} \sum_{i=1}^N x(i), \quad (6)$$

$$d = \left\{ \frac{1}{N-1} \sum_{i=1}^N [x(i) - m]^2 \right\}^{1/2}. \tag{7}$$

Afterwards, we describe the sample $u(i)$ of the random $U(i)$ by a random sequence

$$u(i) = [x(i) - m]/d, \quad i = 1, \dots, N, \tag{8}$$

That is distributed by Tompson's law with $N - 2$ degrees of freedom (Kruopis, 1977):

$$f(N, t) = \begin{cases} \frac{1}{\sqrt{\pi(N-1)}} \frac{\Gamma(\frac{N-1}{2})}{\Gamma(\frac{N-2}{2})} \left(1 - \frac{t^2}{N-1}\right)^{\frac{N-4}{2}}, & |t| < \sqrt{N-1}, \\ 0, & |t| > \sqrt{N-1}, \end{cases} \tag{9}$$

We shall look for the short-time random specific elements $y(i) = x(i) + c(i)$ in the following manner.

Suppose that there are no components $c(i)$ in the sample $y(i)$, $i = 1, \dots, N$, i.e., $c(i) = 0$. Then $y(i) = x(i)$, $i = 1, \dots, N$.

We find the $u(i)$, $i = 1, \dots, N$ element

$$u(k) = \max_{1 \leq i \leq N} |u(i)|. \tag{10}$$

Let us consider the event $A: u(i) \geq u(k)$. A probability for appearance of event A (Kruopis, 1977) is:

$$p = \text{Prob} [U(i) \geq u(k)] = \int_{u(k)}^{\infty} f(N-2, t) dt. \tag{11}$$

By using Bernulli's scheme, we observe the events $A: u(i) \geq u(k)$, $i = 1, \dots, N$. We calculate count the amount quantity n of events $A: u(i) \geq u(k)$, $i = 1, \dots, N$. n of events is distributed by the binomial law (Kruopis, 1977)

$$B(N, p, n) = C_N^n p^n (1-p)^{N-n}. \tag{12}$$

As $n = 1$, the estimate of the probability of event $A: u(i) \geq u(k)$, $i = 1, \dots, N$ is $p(A) = \frac{1}{N}$.

We check hypothesis H that $p(A) < p$ with the reliability level α .

If the data contradict hypothesis H , then we can state with the probability α that $y(k) = x(k) + c(k)$ is a short-time random specific element. In this case, the procedure of search for short-time random specific elements is over.

If the data do not contradict hypothesis H , then the assumption that there are no short-time random specific elements in the observed sequence $y(i)$, $i = 1, \dots, N$ is true with the probability α . It means that there are no short-time random specific elements.

In spite of that, we can proceed with the search of short-time random specific elements by involving more information, making use of differences $y(i) - y(i+r)$, $r = 1, \dots, R$ of the sequence $y(i)$.

Let us define a sequence of random variables

$$Ur(i) = [(X(i) - X(i+r)) - Mr]/Dr, \quad i = 1, 2, \dots, \quad (13)$$

where $Mr = E[X(i) - X(i+r)]$ is the mean of the sequence $X(i) - X(i+r)$, $i = 1, \dots$ and $Dr = [E((X(i) - X(i+r)) - Mr)^2]^{1/2}$ is a standard deviation of the sequence $X(i) - X(i+r)$, $i = 1, \dots$.

Mr and Dr are unknown. We shall use estimates mr and dr instead of Mr and Dr :

$$mr = \frac{1}{N-r} \sum_{i=1}^{N-r} (x(i) - x(i+r)), \quad (14)$$

$$dr = \left\{ \frac{1}{N-r-1} \sum_{i=1}^{N-r} [(x(i) - x(i+r)) - mr]^2 \right\}^{1/2}. \quad (15)$$

Thus, the realizations $ur(i)$ of the random sequence $Ur(i)$ are described by a random sequence

$$ur(i) = \{[x(i) - x(i+r)] - mr\}/dr, \quad (16)$$

That is distributed by the Thompson's law (Kuopis, 1977) with $N - r - 2$ degrees of freedom.

We make an assumption that there are no components $c(i)$ in the sample $y(i)$, $i = 1, \dots, N$, i.e., $c(i) = 0$. Then $y(i) = x(i)$, $i = 1, \dots, N$.

We find an element of $ur(i)$, $i = 1, \dots, N - r$:

$$ur(k) = \max_{1 \leq i \leq N-r} |ur(i)|. \quad (17)$$

Let us deal with the event $Ar: ur(i) \geq ur(k)$. A probability for appearance of the event Ar (Kruopis, 1977) is:

$$pr = \text{Prob}[ur(i) \geq ur(k)] = \int_{ur(k)}^{\infty} f(N-r-2, t) dt. \quad (18)$$

Making use of the Bernulli scheme we observe the events $Ar: ur(i) \geq ur(k)$, $i = 1, \dots, N - r$. We count the quantity n of events $Ar: ur(i) \geq ur(k)$, $i = 1, \dots, N - r$. The quantity n of events is distributed by the binomial law (Kruopis, 1977):

$$B(N-r, pr, nr) = C_{N-r}^{nr} p^{nr} (1-pr)^{N-r-nr}. \quad (19)$$

It the event appears once as $nr = 1$, the estimate of probability of the event Ar is $pr(Ar) = \frac{1}{N-r}$.

We check up hypothesis Hr that $p(Ar) < pr$ with the reliability level α .

If the data do not contradict the hypothesis Hr , then the assumption that there are no short-time random specific elements in the observed sequence $y(i), i = 1, \dots, N$ is true with the probability α . It means that there are no short-time random specific elements in the sequence $y(i), i = 1, \dots, N$ observed. Therefore we continue the search procedure.

If the data contradict hypothesis Hr , then we can affirm with the probability α that $y(k) = x(k) + c(k) \vee y(k+r) = x(k+r) + c(k+r), r = 1, \dots, R$ is a short-time random specific element. In this situation, we terminate the procedure for searching short-time random specific elements.

4. Experimental Investigation

For experimental investigation of the recognition theory created for detecting recognizing short-time specific random elements in random sequences, a software has been developed, which covers 3 MB of computer memory.

We shall of illustrate the results of experimental tests research in search of short-time random specific elements by modelled and real examples of random sequences in recognizing short-time heart rhythm disorders. In all the figures, the vertical segments in their upper part represent values of the sequences $y(i), i = 1, \dots, N$. In their lower part, the some sequences $y(i), i = 1, \dots, N$ are illustrated, where the elements of the sequence marked in black dark are identifiable recognized as short-time random specific elements.

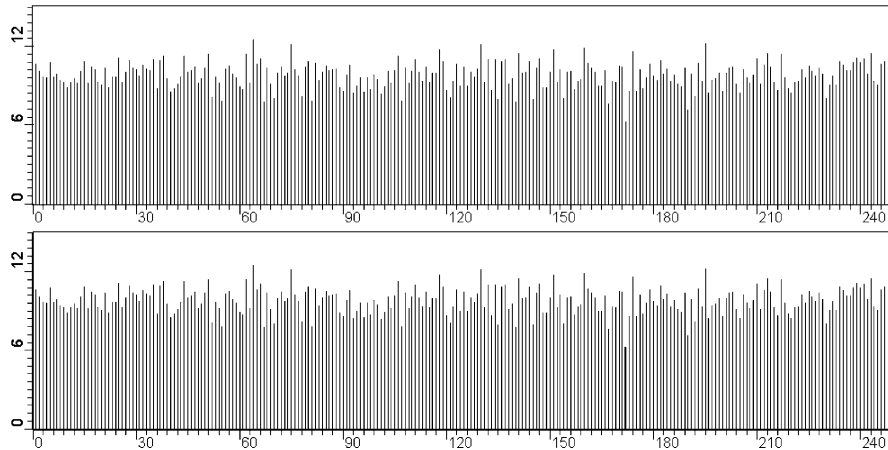


Fig. 1. The upper picture represents the sequence $y(i) = x(i) + 10 + c(i), x(i) = V(i), V(i)$ – Gauss non-correlated random sequence with $EV(i) = 0, EV^2(i) = 1, i = 1, \dots, 250, c(172) = -3.4027$ is illustrated in which there is one short-time random specific element $y(172)$. The lower picture shows that the short-time random specific element $y(172)$ was found correctly.

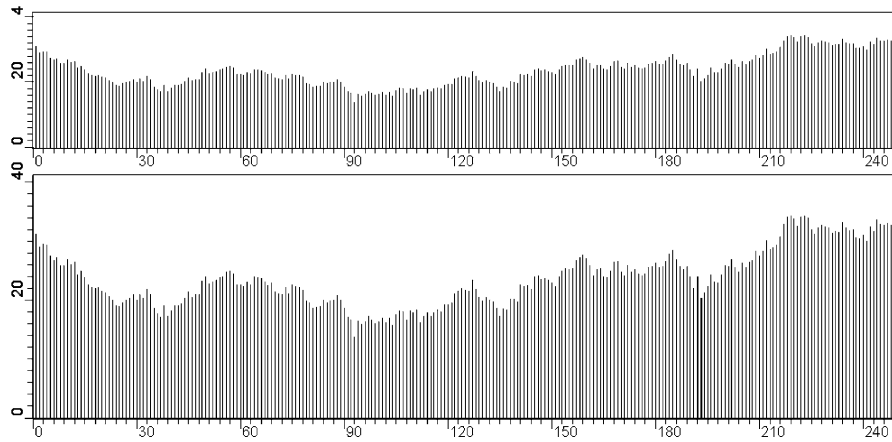


Fig. 2. The upper picture represents the sequence $y(i) = x(i) + 25 + c(i)$, $x(i) - 25 = 0.99[x(i-1) - 25] + V(i)$, $V(i)$ – Gauss non-correlated random sequence with $EV(i) = 0$, $EV^2(i) = 1$, $i = 1, \dots, 250$, $c(192) = 2.4657$ in which there is one short-time random specific element $y(192)$. The lower picture shows short-time specific random elements $y(192) \vee y(193)$, because the decision was made using the difference $y(i) - y(i+1)$, $i = 1, \dots, 249$ of the sequence $y(i)$.

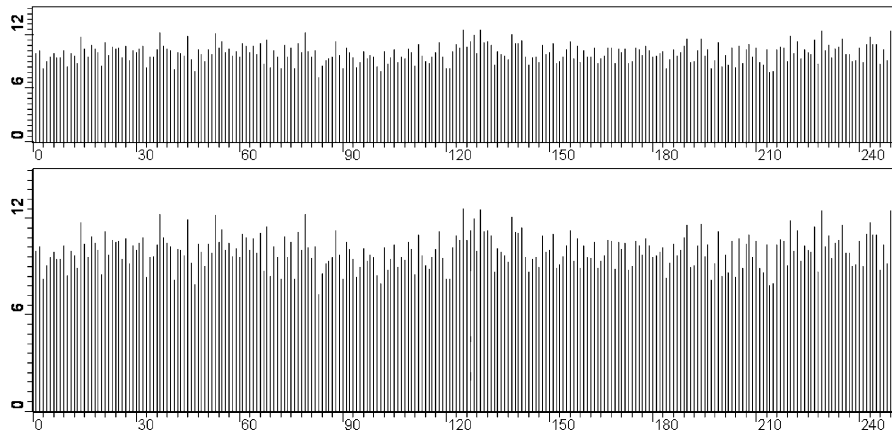


Fig. 3. The upper picture represents the sequence $y(i) = x(i) + 10 + c(i)$, $x(i) = V(i)$, $V(i)$ – Gauss non-correlated random sequence with $EV(i) = 0$, $EV^2(i) = 1$, $i = 1, \dots, 250$, $c(83) = -1.9872$. is a short-time random specific element $y(83)$. The lower figure shows that a short-time random specific elements was not-found.

The modelled results when $\alpha = 95\%$ are presented in random sequences in which there are:

- are short-time random specific elements and they are found (Figs. 1 and 2);
- are short-time random specific elements, but they are not found (Fig. 3);
- are no short-time random specific elements, but they are found (Fig. 4);
- are no short-time random specific elements and they are not found (Fig. 5).

We give two examples that illustrate the usage of the method for actual solution of problems. We observe action of the heart of a freely moving person, using the computer equipment in wear. We register time intervals between adjacent systoles and obtain a random sequence $y(i)$, $i = 1, \dots, N$ called a rhythmogram. Making use of the given method, in the rhythmogram we shall look for short-time random specific elements that emerge due to heart rhythm (Fig. 7) or conduction disturbances (Fig. 6).

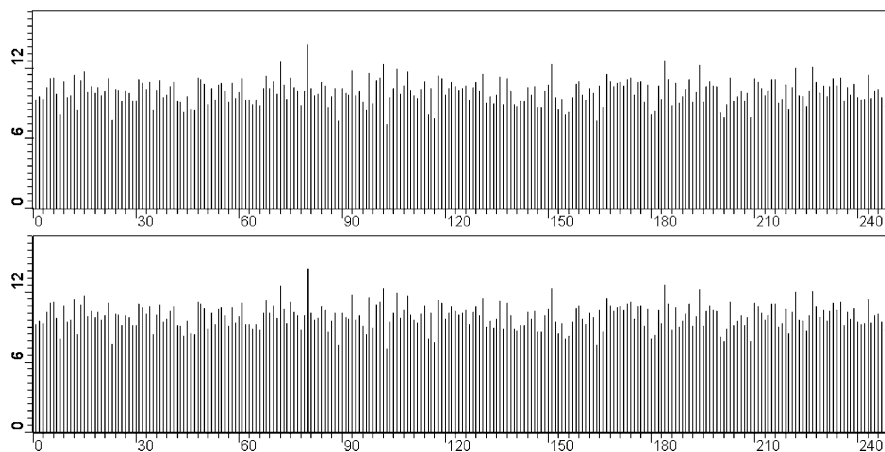


Fig. 4. The upper picture represents the sequence $y(i) = x(i) + 10$, $x(i) = V(i)$, $V(i)$ – Gauss non-correlated random sequence with $EV(i) = 0$, $EV^2(i) = 1$, $i = 1, \dots, 250$ in which there are no short-time random specific elements. In the lower picture, an element of the sequence is marked in black dark which was faulty recognized as a short-time random specific element.

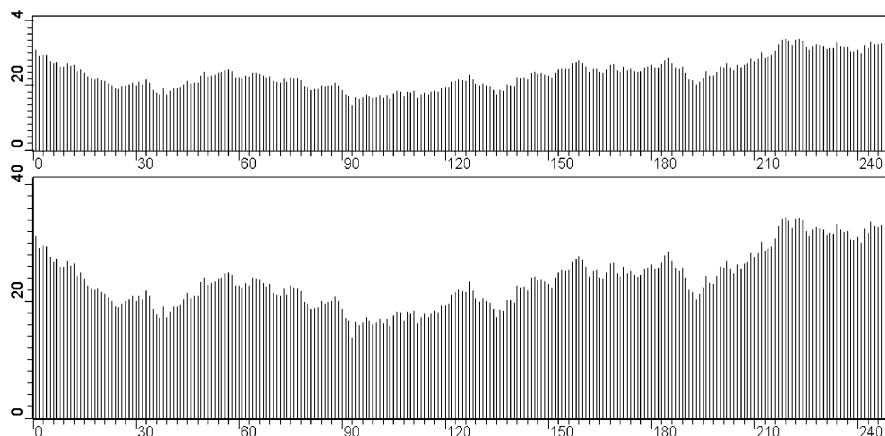


Fig. 5. The upper picture represents the sequence $y(i) = x(i) + 25$, $x(i) - 25 = 0.99[x(i-1) - 25] + V(i)$, $V(i)$ – Gauss non-correlated random sequence with $EV(i) = 0$, $EV^2(i) = 1$, in which there are no short-time random specific elements. The lower picture shows that the recognition system has made correct decisions: short-time random specific elements are not found.

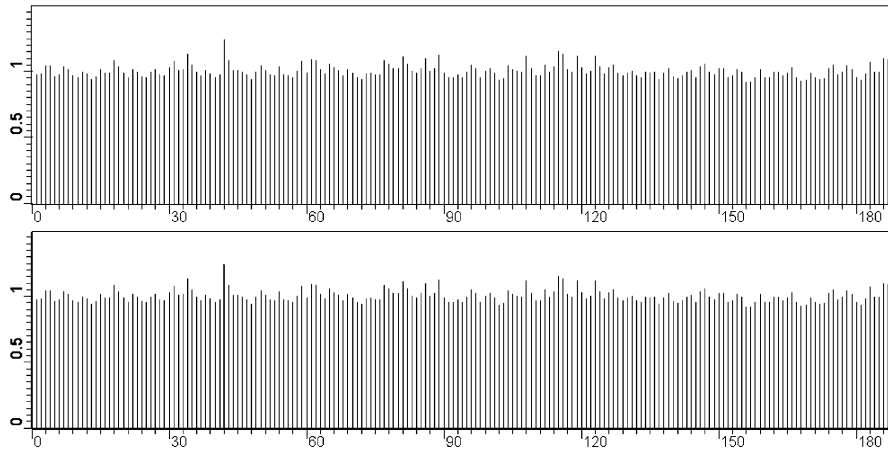


Fig. 6. The upper picture represents the sequence $y(i)$, $i = 1, \dots, 188$ in which there is a conduction trouble. The lower picture shows that the recognition system has made a correct decision.

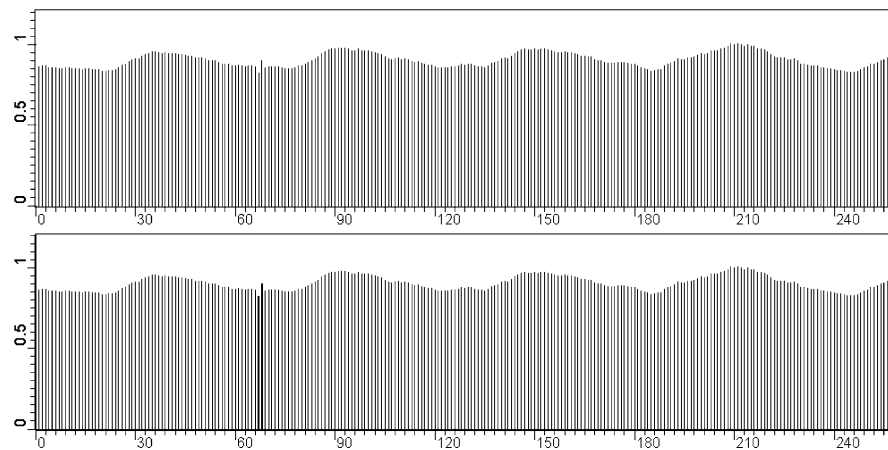


Fig. 7. The upper picture represents the sequence $y(i)$, $i = 1, \dots, 258$ in which there is a premature beats. The lower picture shows that the recognition system has made a correct decision.

5. Conclusions

The theory and the constructive method presented provides an opportunity to recognize single short-time specific random elements that appear in the background of random sequences.

The probability that decisions about the existence of an isolated short-time specific random element in the background of a random sequence, while it doesn't exist in reality, will be made, is posed and controlled on the reliability level α of hypothesis verification.

There are no possibilities to estimate the probability of individual short-time specific elements missed without their statistical characteristics. However, this shortcoming can

be eliminated. To this end, it is reasonable to invoke the theory and method for recognizing short-time random specific elements, described in this paper. They provide with a possibility to accumulate missing information by analyzing the properties of short-time specific random elements and use it in the estimation of probabilities of short-time specific random elements.

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Trumpalaikių specifinių atsitiktinių elementų atpažinimas atsitiktinėse sekose

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Nagrinėjamos atsitiktinės sekos, kurių fone atsitiktiniais laiko momentais gali pasirodyti trumpalaikiai atsitiktiniai specifiniai sekos elementai. Išdėstyta teorija ir konstruktyvūs metodai trumpalaikių atsitiktinių specifinių elementų, galinčių pasirodyti atsitiktinių sekų fone atpažinimui. Pateikiami eksperimentinių tyrimų rezultatai. Aptartos gautų rezultatų platesnio panaudojimo perspektyvos.